

rags2ridges  
Network-Based Modeling and Analysis of  
High-Dimensional Omics Data

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Bioinformatics Seminar CCA  
Amsterdam, the Netherlands  
March 10, 2017



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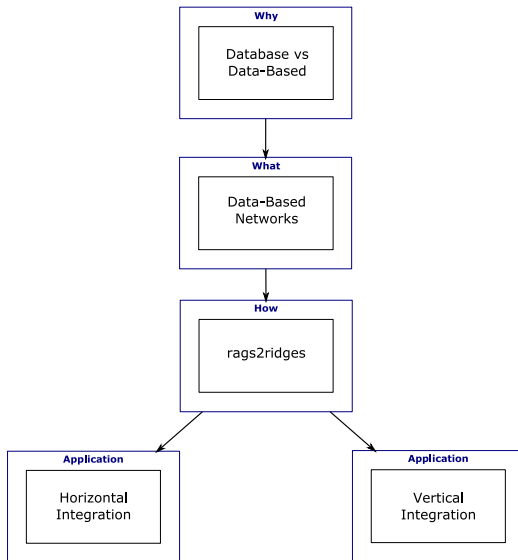


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# Pathway Databases

## Metabolic pathway databases

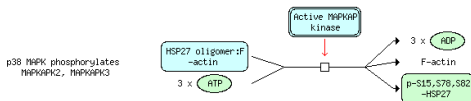
- MetaCyc
- HumanCyc
- BioCyc

## Signaling pathway databases

- KEGG
- BioCarta
- Reactome
- Biomodels
- Human signaling network
- Ingenuity
- PID
- BioPP

## Manual Curation

## 1.2.50 p-MAPK2/3 phosphorylates HSP27 (BlackBoxEvent)

**Authors**

Garapati, P V, 2013-08-30.

**Reviewers**

Welsh, Michael, Berger, Philipp, Ballmer-Hofer, Kurt, 2014-05-12.

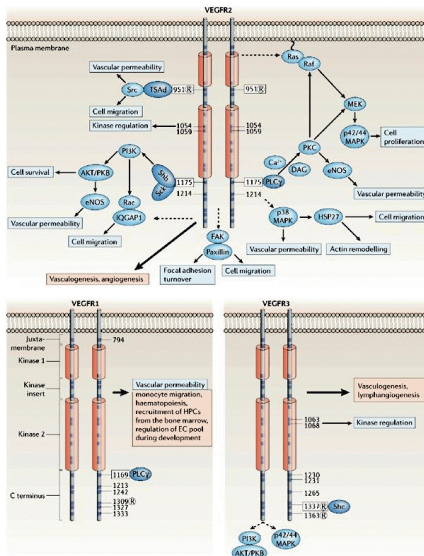
Activated MAP kinase-activated protein kinase (MAPK/MAPKAPK) 2 and 3 in turn phosphorylate heat shock protein beta 1 (HSPB1, HSP27).

**References**

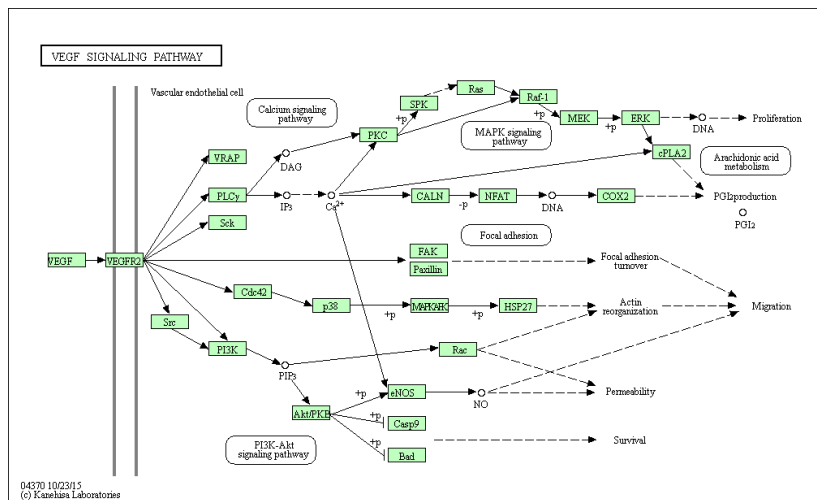
Landry J, Huot J, "Regulation of actin dynamics by stress-activated protein kinase 2 (SAPK2)-dependent phosphorylation of heat-shock protein of 27 kDa (Hsp27)", *Biochem. Soc. Symp.*, 64, 1999, 79-89.

Lavoie JN, Lambert H, Hickey E, Weber LA, Landry J, "Modulation of cellular thermoresistance and actin filament stability accompanies phosphorylation-induced changes in the oligomeric structure of heat shock protein 27", *Mol. Cell. Biol.*, 15, 1995, 505-16.

## VEGF Signaling According to REACTOME

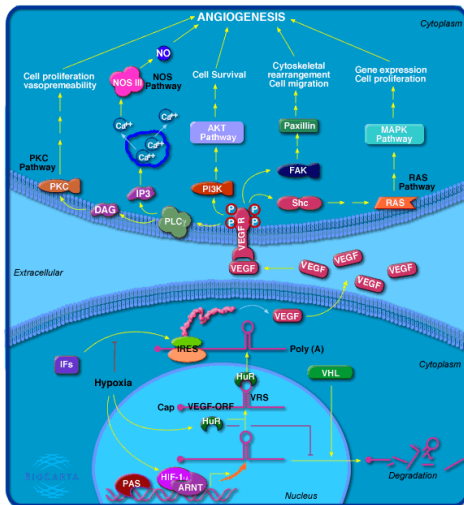


## VEGF Signaling According to KEGG





## VEGF Signaling According to BIOCARTA



# Graphs

## Representation

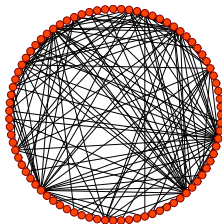
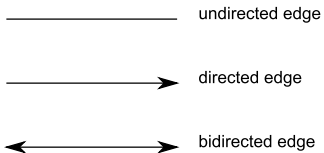
Pathways are represented by a *graph* (or *network*)

## Vertices

○ *Node* or *vertex* represents molecular feature

## Edges

*Edge* or *arrow* represents some functional relation



# Correlation networks

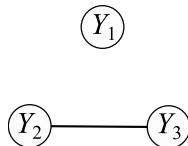
## Example

Three variables:  $Y_1$ ,  $Y_2$ , and  $Y_3$

$$\text{cor}(Y_1, Y_2) = 0$$

$$\text{cor}(Y_1, Y_3) = 0$$

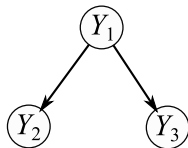
$$\text{cor}(Y_2, Y_3) \neq 0$$



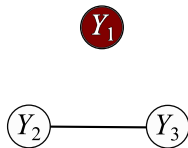
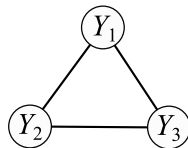
## Marginal dependence

Undirected edge represents marginal dependence

## Interpretational danger



True mechanism

Not observing  $Y_1$ :  
Spurious associationObserving  $Y_1$ :  
Saturated graph

## Conditional dependence

### Partial correlation

Measures degree of association between two random variables when controlling for third variables

### Conditioned correlation

$$\text{cor}(Y_1, Y_2|Y_3)$$

$$\text{cor}(Y_1, Y_3|Y_2)$$

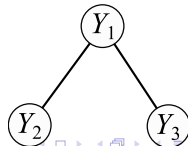
$$\text{cor}(Y_2, Y_3|Y_1)$$

If, e.g.,  $\text{cor}(Y_2, Y_3|Y_1) = 0$ , we say  $Y_2$  and  $Y_3$  are independent *given*  $Y_1$

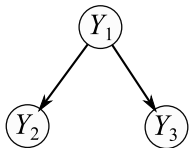
$$\text{cor}(Y_1, Y_2|Y_3) \neq 0$$

$$\text{cor}(Y_1, Y_3|Y_2) \neq 0$$

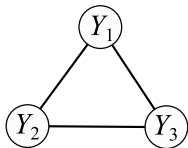
$$\text{cor}(Y_2, Y_3|Y_1) = 0$$



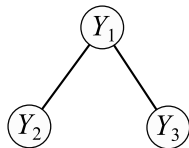
# Partial Correlation Networks



True mechanism



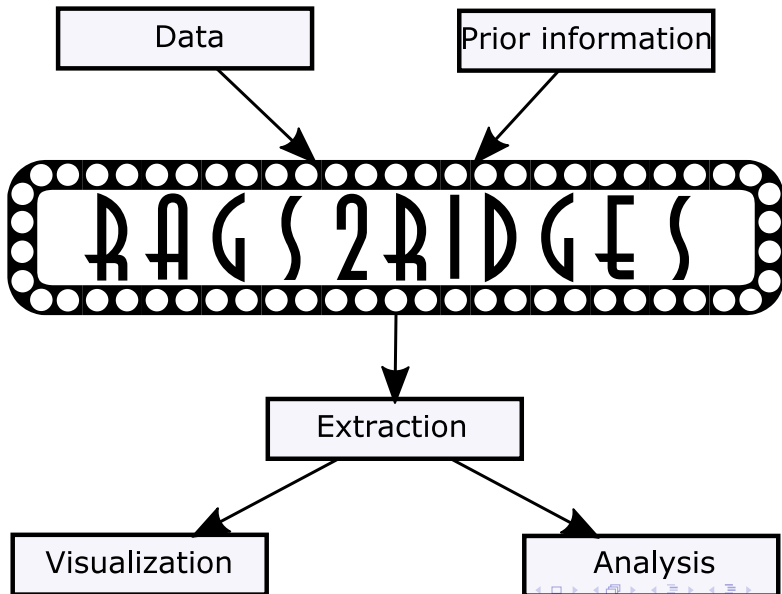
Correlation graph



Conditional independence graph

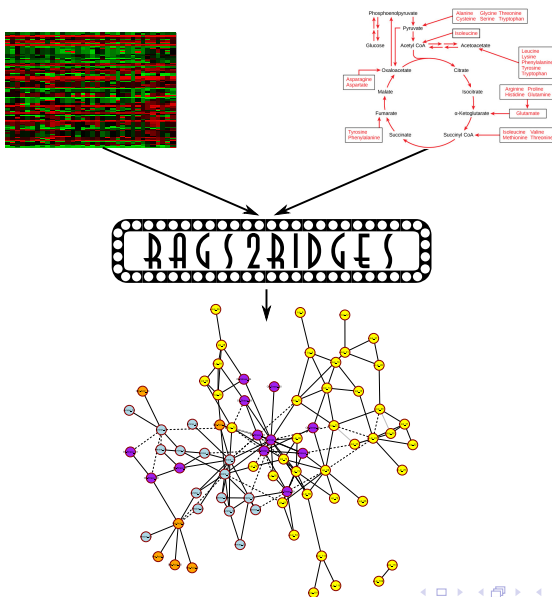


## One-Stop-Go

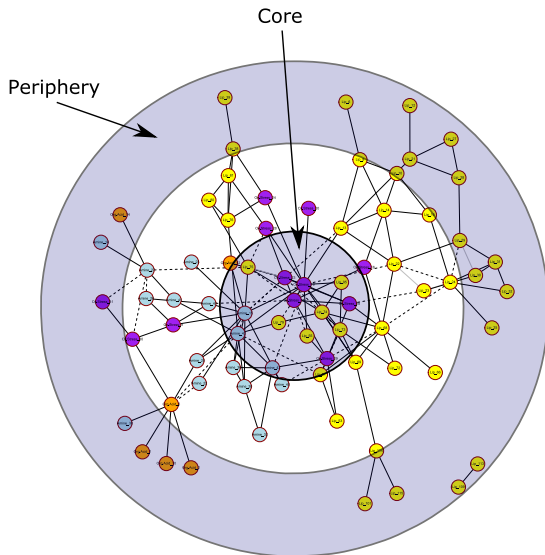




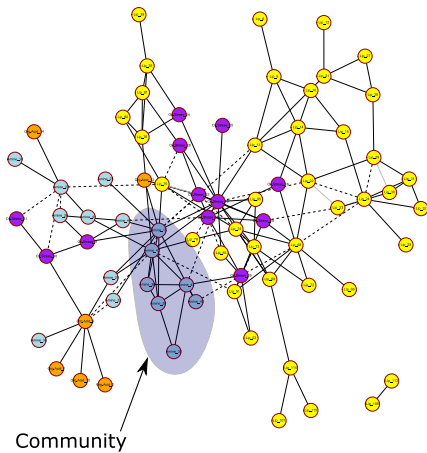
# One-Stop-Go



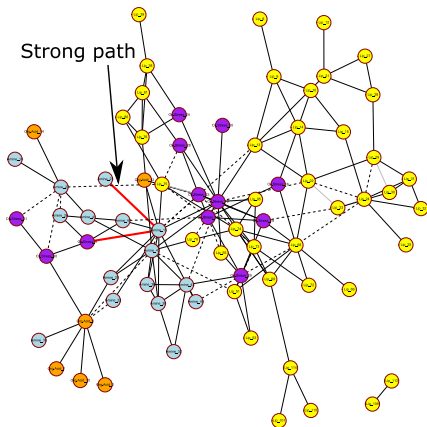
## Analysis: Global Level



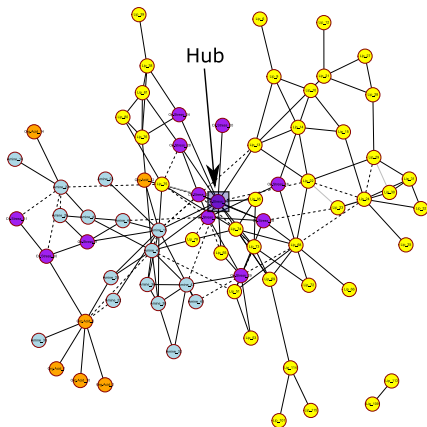
# Analysis: Group Level



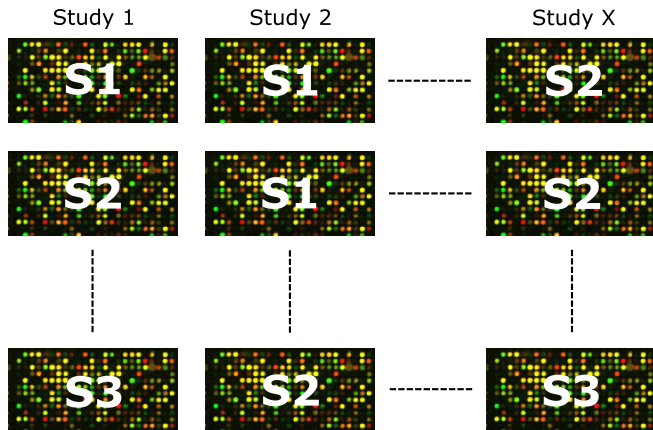
## Analysis: Path Level



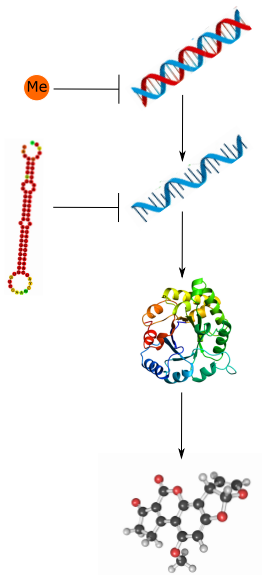
## Analysis: Node Level



## Integration: Horizontal



## Integration: Vertical



# Background

## DLBCL

- Diffuse large B-cell lymphomas
- A non-Hodgkin type of blood cancer

## DLBCL subtypes

At least two major genetic subtypes of tumors:

- ABC: activated B-cells
- GCB: germinal centre B-cells
- III: cannot be classified as either ABC or GCB

## NF- $\kappa$ B signaling pathway

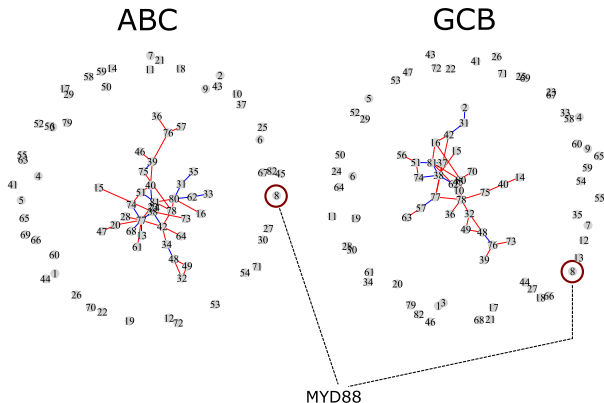
- Responsible for control of cell survival
- Known to be deregulated in DLBCL
- Hallmark distinguishing *poor prognostic* ABC from *good prognostic* GCB



# A non-integrative analysis

## Data

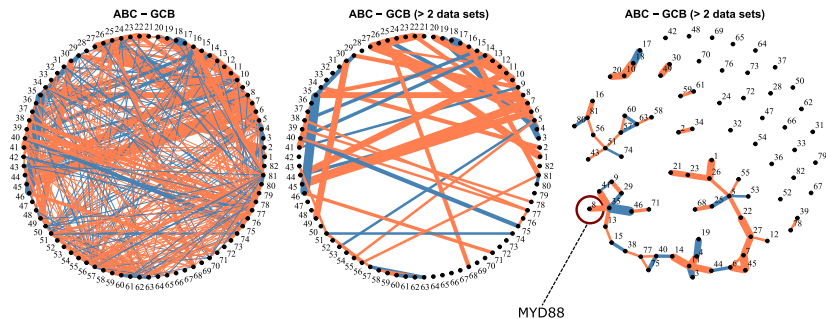
- $n = 89$  DLBCL tumor samples
- ABC ( $n_1 = 31$ ), III ( $n_2 = 13$ ), and GCB ( $n_3 = 45$ )
- $p = 82$  (KEGG)



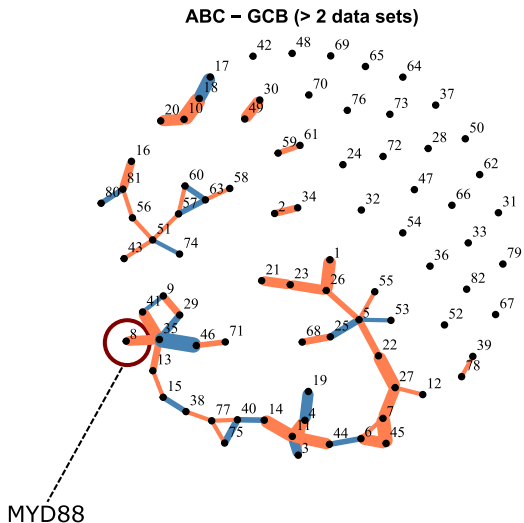
## Integrative analysis: Data

	ABC		Type III		GBC		$\sum n_g$
	$g$	$n_g$	$g$	$n_g$	$g$	$n_g$	
<b>Pilot data</b>							
GSE11318		74		71		27	172
<b>Data set</b>							
GSE56315	1	31	2	13	3	45	89
GSE19246	4	51	5	30	6	96	177
GSE12195	7	40	8	18	9	78	136
GSE22895	10	31	11	21	12	49	101
GSE31312	13	146	14	97	15	224	467
GSE10846.CHOP	16	64	17	28	18	89	181
GSE10846.RCHOP	19	75	20	42	21	116	233
GSE34171.hgu133plus2	22	23	23	15	24	52	90
GSE34171.hgu133AplusB	25	18	26	17	27	43	78
GSE22470	28	86	29	43	30	142	271
GSE4475	31	73	32	20	33	128	221
$\sum n_g$		638		344		1062	2044

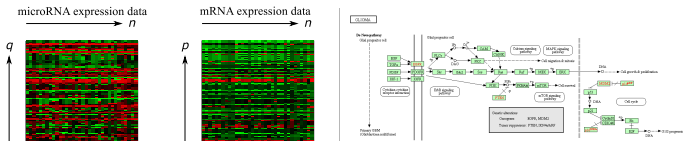
## Integrative analysis: Results



## Integrative analysis: Results

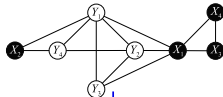


## Background

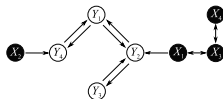


1. Regularized estimation of joint precision matrix

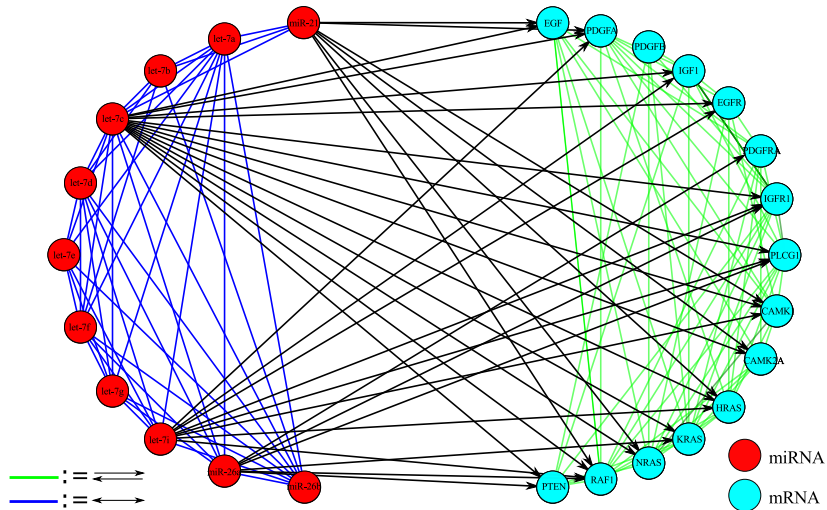
2. Determine support. Resulting sparsified matrix represents the moral graph



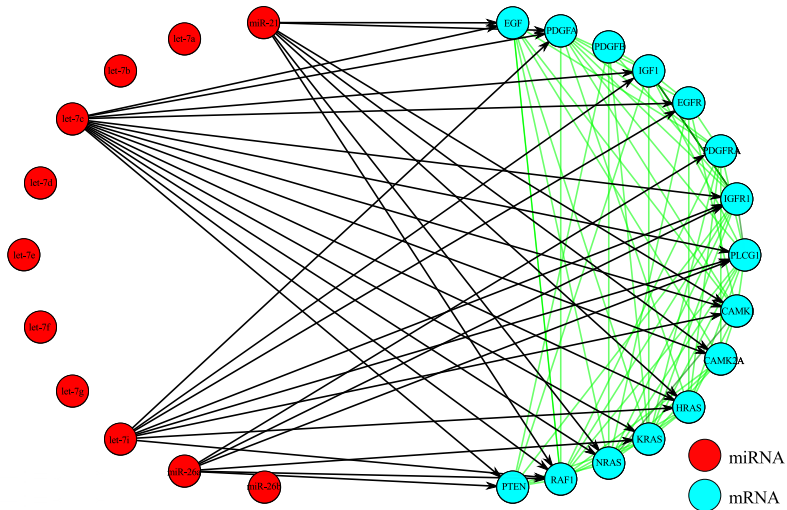
3. Run decomposition algorithm to obtain the cyclic mixed graph



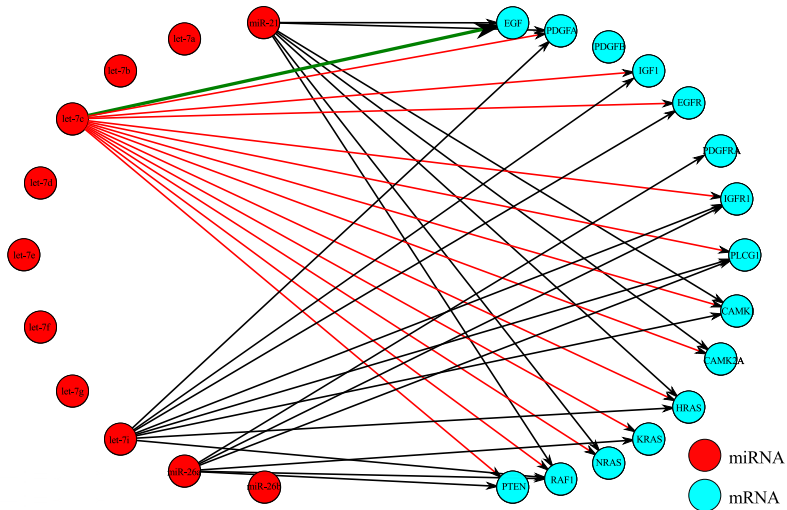
## Full DCMG



## Endogenous Relations



## Exogenous Shocks





## Manual/Download

- Peeters, C.F.W., Bilgrau, A.E., & van Wieringen, W.N. (2016). “rags2ridges: Ridge Estimation of Precision Matrices from High-Dimensional Data”. R package, version 2.1.1 URL: <https://cran.r-project.org/package=rags2ridges>.

## Theory/Methodology

- Bilgrau\*, A.E., Peeters\*, C.F.W., Eriksen, P.S., Bøgsted, M., & van Wieringen, W.N. (2015). “Targeted Fused Ridge Estimation of Inverse Covariance Matrices from Multiple High-Dimensional Data Classes”. [arXiv:1509.07982v1](https://arxiv.org/abs/1509.07982v1) [stat.ME].
- Peeters, C.F.W., van Wieringen, W.N., & van de Wiel, M.A. (in preparation). “Directed Cyclic Mixed Graph Modeling for High-Dimensional Genomic Data Integration”.
- van Wieringen, W.N. & Peeters, C.F.W. (2016). “Ridge Estimation of Inverse Covariance Matrices from High-Dimensional Data”. *Computational Statistics & Data Analysis*, 103: 284-303. [arXiv:1403.0904v3](https://arxiv.org/abs/1403.0904v3) [stat.ME].

## Software

- Peeters, C.F.W., van de Wiel, M.A., & van Wieringen, W.N. (2016) “The Spectral Condition Number Plot for Regularization Parameter Determination”. [arXiv:1608.04123v1](https://arxiv.org/abs/1608.04123v1) [stat.CO].
- van Wieringen, W.N. & Peeters, C.F.W. (2015). “Application of a New Ridge Estimator of the Inverse Covariance Matrix to the Reconstruction of Gene-Gene Interaction Networks”. In: di Serio, C., Lio, P., Nonis, A., and Tagliaferri, R. (Eds.) ‘Computational Intelligence Methods for Bioinformatics and Biostatistics’. *Lecture Notes in Computer Science*, vol. 8623. Springer, pp. 170–179.

## Explaining the inverse

### The scalar inverse

- Let  $a$  denote a number (excluding 0)
- The inverse is then the number  $b$  such that  $a \times b = 1$
- Clearly,  $b = \frac{1}{a}$

### Matrix

A matrix is a generalization of a number, an array of numbers

$$\mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1p} \\ a_{21} & a_{22} & \cdots & a_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ a_{p1} & a_{p2} & \cdots & a_{pp} \end{bmatrix}$$

# Explaining the inverse

## The Matrix Inverse

Consider the matrix  $\mathbf{A}$ . Its inverse  $\mathbf{B} = \mathbf{A}^{-1}$  is defined such that

$$\mathbf{AB} = \mathbf{I},$$

where

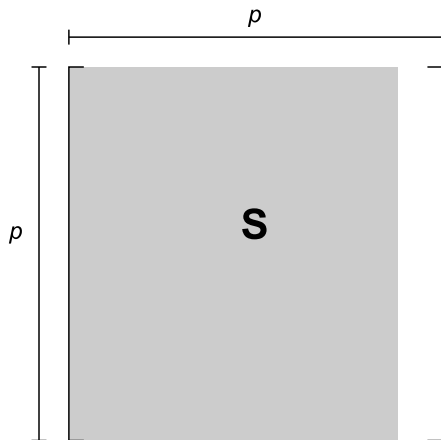
$$\mathbf{I} = \begin{bmatrix} 1 & 0 & \cdots & 0 \\ 0 & 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1 \end{bmatrix}$$

## Solution

$$\mathbf{A}^{-1} = \begin{bmatrix} \mathbf{A}_{11}^{-1} + \mathbf{A}_{11}^{-1} \mathbf{A}_{12} \mathbf{Q}^{-1} \mathbf{A}_{21} \mathbf{A}_{11}^{-1} & -\mathbf{A}_{11}^{-1} \mathbf{A}_{12} \mathbf{Q}^{-1} \\ -\mathbf{Q}^{-1} \mathbf{A}_{21} \mathbf{A}_{11}^{-1} & \mathbf{Q}^{-1} \end{bmatrix},$$

with  $\mathbf{Q} = \mathbf{A}_{22} - \mathbf{A}_{21} \mathbf{A}_{11}^{-1} \mathbf{A}_{12}$  denoting the Schur complement.

# Singularity



# Ridge estimation

Maximize

$$\underbrace{\ln |\mathbf{\Omega}| - \text{tr}(\mathbf{S}\mathbf{\Omega})}_{\text{log-likelihood}} - \underbrace{\frac{\lambda}{2} \|\mathbf{\Omega} - \mathbf{T}\|_2^2}_{\ell_2\text{-penalty}}$$

- $\mathbf{T}$  denotes a p.d. symmetric target matrix
- $\lambda \in (0, \infty)$  denotes a penalty parameter

Analytic penalized ML estimator

$$\hat{\mathbf{\Omega}}(\lambda) = \left\{ \left[ \lambda \mathbf{I}_p + \frac{1}{4}(\mathbf{S} - \lambda \mathbf{T})^2 \right]^{1/2} + \frac{1}{2}(\mathbf{S} - \lambda \mathbf{T}) \right\}^{-1}$$

# Properties

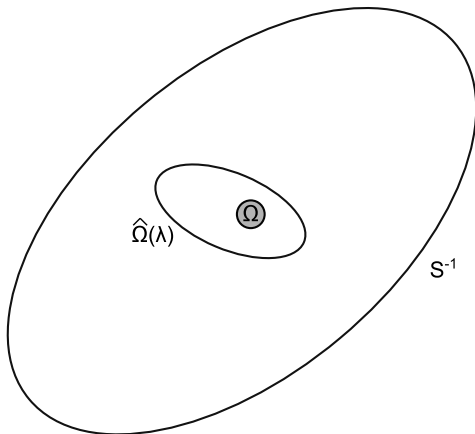
## Behavior

- i.  $\hat{\Omega}(\lambda) \succ 0$ , for all  $\lambda \in (0, \infty)$ ;
- ii.  $\lim_{\lambda \rightarrow 0^+} \hat{\Omega}(\lambda) = \mathbf{S}^{-1}$  if  $p < n$ ;
- iii.  $\lim_{\lambda \rightarrow \infty} \hat{\Omega}(\lambda) = \mathbf{T}$ .

## Consistency

- i.  $\lim_{n \rightarrow \infty} \mathbb{E} \left[ \hat{\Omega}_n(\lambda_n) \right] \longrightarrow \lim_{n \rightarrow \infty} \mathbb{E} \left( \mathbf{S}_n^{-1} \right) = \Omega$ ;
- ii.  $\lim_{n \rightarrow \infty} \mathbb{E} \left( \left\| \hat{\Omega}_n(\lambda_n) - \Omega \right\|_F^2 \right) = 0$ .

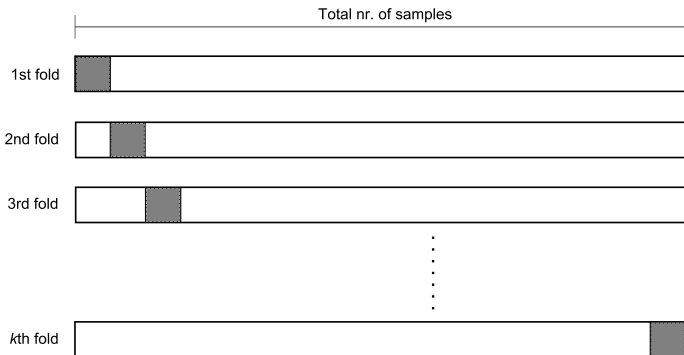
## Visual explanation



# Choosing the penalty value

## $K$ -fold cross-validation (CV)

### Single iteration of $K$ -fold CV



Test data



Training data



## Choosing the penalty value

### $K$ -fold CV score

$$\varphi^K(\lambda) = \sum_{k=1}^K n_k \left\{ -\ln |\hat{\Omega}(\lambda)_{-k}| + \text{tr}[\hat{\Omega}(\lambda)_{-k} \mathbf{S}_k] \right\},$$

$n_k$  is the size of subset  $k$ , for  $k = 1, \dots, K$  disjoint subsets;

$\mathbf{S}_k$  denotes the sample covariance matrix on  $k$ th test set;

$\hat{\Omega}(\lambda)_{-k}$  denotes the estimated regularized precision matrix on  $k$ th training set

### Highest predictive accuracy

Choose  $n_k = 1$ , such that  $K = n$  (known as leave-one-out CV - LOOCV)

### Problem

$K$ -fold CV is computationally demanding for large  $p$  and/or large  $K$

### Solution

Computationally efficient approximate LOOCV score

## Support determination

### Scaling

$\hat{\mathbf{P}}(\lambda)$ : Regularized precision estimate scaled to partial correlation form

### Assume

Nonredundant off-diagonal partial correlation coefficients (indexed by  $j < j'$ ) follow a mixture distribution:

$$f \left\{ [\hat{\mathbf{P}}(\lambda^*)]_{jj'} \right\} = \eta_0 f_0 \left\{ [\hat{\mathbf{P}}(\lambda^*)]_{jj'}; \kappa \right\} + (1 - \eta_0) f_{\mathcal{E}} \left\{ [\hat{\mathbf{P}}(\lambda^*)]_{jj'} \right\}$$

- $\eta_0 \in [0, 1]$  is the mixture weight
- $f_0 \{ \cdot \}$  denotes the distribution of a null-edge
- $f_{\mathcal{E}} \{ \cdot \}$  denotes the distribution of a present edge

### Determine

$$P \left( Y_j \neq Y_{j'} \mid [\hat{\mathbf{P}}(\lambda^*)]_{jj'} \right)$$

# Situation

## Data

- $G$  classes of  $(n_g \times p)$ -dimensional data
- Classes defined by data sets and/or (subtypes of) diseases

## Assumption

Precision matrices of constituent classes chiefly share the same structure but potentially differ in a number of locations of interest

## Desire

Integrative or meta-analytic Gaussian graphical modeling

## Targeted fused ridge estimation: General Formulation

Maximize

$$\underbrace{\mathcal{L}(\{\Omega_g\}; \{S_g\})}_{\text{log-likelihood}} - \sum_g \underbrace{\frac{\lambda_{gg}}{2} \|\Omega_g - T_g\|_F^2}_{\text{ridge-penalty}} - \sum_{g_1, g_2} \underbrace{\frac{\lambda_{g_1 g_2}}{4} \|(\Omega_{g_1} - T_{g_1}) - (\Omega_{g_2} - T_{g_2})\|_F^2}_{\text{fusion-penalty}}$$

- $T_g$  indicate class-specific target matrices
- $\lambda_{gg} \in (0, \infty)$  denote class-specific ridge penalty parameters
- $\lambda_{g_1 g_2} \in [0, \infty)$  denote pair-specific fusion penalty parameters,  $\lambda_{g_1 g_2} = \lambda_{g_2 g_1}$

Penalty matrix

All penalties can be collected into a non-negative symmetric matrix  $\Lambda = [\lambda_{g_1 g_2}]$

## Targeted fused ridge estimation

Maximizing argument for class  $g_0$

$$\hat{\Omega}_{g_0}(\Lambda, \{\Omega_g\}_{g \neq g_0}) = \left\{ \left[ \bar{\lambda}_{g_0} \mathbf{I}_p + \frac{1}{4} (\bar{\mathbf{S}}_{g_0} - \bar{\lambda}_{g_0} \mathbf{T}_{g_0})^2 \right]^{1/2} + \frac{1}{2} (\bar{\mathbf{S}}_{g_0} - \bar{\lambda}_{g_0} \mathbf{T}_{g_0}) \right\}^{-1},$$

where

$$\bar{\mathbf{S}}_{g_0} = \mathbf{S}_{g_0} - \sum_{g \neq g_0} \frac{\lambda_{gg_0}}{n_{g_0}} (\Omega_g - \mathbf{T}_g), \quad \text{and} \quad \bar{\lambda}_{g_0} = \frac{\sum_g \lambda_{gg_0}}{n_{g_0}}$$

# Properties

## Behavior

- i.  $\hat{\Omega}_g \succ \mathbf{0}$  for all  $\lambda_{gg} \in (0, \infty)$ ;
- ii.  $\lim_{\lambda_{gg} \rightarrow 0^+} \hat{\Omega}_g = \mathbf{S}_g^{-1}$  if  $\sum_{g' \neq g} \lambda_{gg'} = 0$  and  $p \leq n_g$ ;
- iii.  $\lim_{\lambda_{gg} \rightarrow \infty} \hat{\Omega}_g = \mathbf{T}_g$  if  $\lambda_{gg'} < \infty$  for all  $g' \neq g$ ;
- iv.  $\lim_{\lambda_{g_1 g_2} \rightarrow \infty} (\hat{\Omega}_{g_1} - \mathbf{T}_{g_1}) = \lim_{\lambda_{g_1 g_2} \rightarrow \infty} (\hat{\Omega}_{g_2} - \mathbf{T}_{g_2})$  if  $\lambda_{g'_1 g'_2} < \infty$  for all  $\{g'_1, g'_2\} \neq \{g_1, g_2\}$ .

## Block coordinate ascent

- 1: **Input:**
- 2: *Sufficient data:*  $(\mathbf{S}_1, n_1), \dots, (\mathbf{S}_G, n_G)$
- 3: *Penalty matrix:*  $\mathbf{\Lambda}$
- 4: *Convergence criterion:*  $\varepsilon > 0$
- 5: **Output:**
- 6: *Estimates:*  $\hat{\mathbf{\Omega}}_1, \dots, \hat{\mathbf{\Omega}}_G$
- 7: **procedure** RIDGEP.FUSED( $\mathbf{S}_1, \dots, \mathbf{S}_G, n_1, \dots, n_G, \mathbf{\Lambda}, \varepsilon$ )
- 8:     *Initialize:*  $\hat{\mathbf{\Omega}}_g^{(0)}$  for all  $g$ .
- 9:     **for**  $c = 1, 2, 3, \dots$  **do**
- 10:         **for**  $g = 1, 2, \dots, G$  **do**
- 11:             Update  $\hat{\mathbf{\Omega}}_g^{(c)} := \hat{\mathbf{\Omega}}_g(\mathbf{\Lambda}, \hat{\mathbf{\Omega}}_1^{(c)}, \dots, \hat{\mathbf{\Omega}}_{g-1}^{(c)}, \hat{\mathbf{\Omega}}_{g+1}^{(c-1)}, \dots, \hat{\mathbf{\Omega}}_G^{(c-1)})$
- 12:         **end for**
- 13:         **if**  $\max_g \left\{ \frac{\|\hat{\mathbf{\Omega}}_g^{(c)} - \hat{\mathbf{\Omega}}_g^{(c-1)}\|_F^2}{\|\hat{\mathbf{\Omega}}_g^{(c)}\|_F^2} \right\} < \varepsilon$  **then**
- 14:             **return**  $(\hat{\mathbf{\Omega}}_1^{(c)}, \dots, \hat{\mathbf{\Omega}}_G^{(c)})$
- 15:         **end if**
- 16:     **end for**
- 17: **end procedure**